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APPROXIMATING CONDITIONAL MOMENTS OF THE MULTIVARIATE NORMAL DISTRIBUTION

BY

JOSEPH G. DEKEN

TECHNICAL REPORT NO. 41
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OFFICE OF NAVAL RESEARCH

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DEPARTMENT OF STATISTICS STANFORD UNIVERSITY STANFORD, CALIFORNIA





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# APPROXIMATING CONDITIONAL MOMENTS OF THE MULTIVARIATE

### NORMAL DISTRIBUTION

By

Joseph G. Deken

## **ABSTRACT**

A practical method for computing the conditional expectation of a polynomial in the components of a multivariate normal random variable X, when X is restricted to a subset of  $\mathbb{R}^p$ , is given. This method makes the application of certain missing data techniques possible in cases where repeated numerical integration is not feasible.

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Key words: Multiple integration, numerical integration, multivariate normal distribution, FM estimation.

# APPROXIMATING CONDITIONAL MOMENTS OF THE MULTIVARIATE NORMAL DISTRIBUTION

Ву

#### Joseph G. Deken

### 1. Introduction.

The conditional moments such as  $E(X_j^k)$  of a multivariate normal random variable  $X = (X_1, X_2, \dots, X_p)$ , when X is restricted to a subset  $A \subset \mathbb{R}^p$ , are not readily obtained numerically, since the required integration in p-dimensions is time-consuming except for very small p. These conditional moments are of interest, for example in the derivation of E-M estimates in missing data problems (Dempster, Laird and Rubin, 1977). We present here an efficient approximation scheme for these moments, which makes the computation practical for moderately large p.

For convenience of description, we restrict attention to sets A of the form  $I_1 \times I_2 \times \cdots \times I_p$ , where all the  $I_j$  are intervals, but the approach is more general, as indicated below. We start by observing that the approximations to the ratio

$$E_{k} = \frac{s+t \int x^{k} e^{-\frac{(x-\mu)^{2}}{2\sigma^{2}}} dx}{s+t \int e^{-\frac{(x-\mu)^{2}}{2\sigma^{2}}} dx}$$

obtained by the first terms in a Taylor series around t = 0 may be written as polynomials in u:

$$E_k \approx \sum_{\alpha=0}^{m} c_{k\alpha}(s,t,\sigma^2)_{\mu}^{\alpha}$$
.

The conditional expectation of a polynomial  $P(X) = a_0 + a_1 X + a_2 X^2 + \cdots + a_n X^n$  in the normal random variable X is thus approximated by replacing P by a polynomial  $Q(\mu)$  in the mean of X, where the transformation  $T(s,t,\sigma^2)$ :  $P \to Q$  is linear. Since the transformation V:  $c \to b$  defined by

$$\sum_{\alpha=0}^{m} c_{\alpha}(g+hX)^{\alpha} = \sum_{\beta=0}^{m} b_{\beta}X^{\beta}$$

is also linear, and the conditional mean of  $X_p$  given  $X_1,\dots,X_{p-1}$  is of the form  $(g+hX_{p-1})$  where g is linear in  $X_1,\dots,X_{p-2}$ , any approximation to  $EX_p^k$  which is a polynomial in the mean of  $X_p$  will produce a polynomial in  $(X_1,\dots,X_{p-1})$  when  $X_p$  is conditional on  $X_1,\dots,X_{p-1}$ . The conditional expectation in  $\mathbb{R}^p$  may thus be accomplished in 2p steps, by applying T and then V in succession p times, to obtain

$$E(\cdot|\mathbf{X}_1\cdots\mathbf{X}_{p-1}) \approx E(\cdot|\mathbf{X}_1\cdots\mathbf{X}_{p-2}) \approx \cdots \approx E(\cdot|\mathbf{X}_1) \approx \sum_{\alpha=0}^{N} g_{\alpha}^{\alpha}.$$

Computationally, this process requires 2p matrix multiplications and is thus practical for moderate p.

# 2. An Example.

The following simple example will serve to establish some ideas and notation. Let

$$I_{k}(\sigma^{2},\mu,s,t) = \int_{s-t}^{s+t} x^{k} e^{-\frac{(\mu-x)^{2}}{2\sigma^{2}}} dx.$$

We are concerned with the conditional expectation  $I_k/I_0$  of  $X^k$ , where X is a normal random variable. The sum of the first terms of a Taylor series for this ratio about t=0 is of the form

(1) 
$$\sum_{\alpha=0}^{\mathbf{q}} c_{\mathbf{k}\alpha}^{*}(\sigma^{2}, \mu, s) \mathbf{t}^{2\alpha} ,$$

but also may be represented by identifying the coefficients of  $\mu^{\alpha}$  in (1) as

$$\sum_{\alpha=0}^{2\mathbf{q}} c_{\alpha}(\sigma^2, s, t)_{\mu}^{\alpha}.$$

For example, (letting  $v = \sigma^2$ )

$$I_{1}/I_{0} = s + \frac{(\mu - s)t^{2}}{3v} + \frac{2v(s - \mu) + (s - \mu)^{3}}{45v^{3}}t^{\frac{1}{4}} + \dots$$

$$= s(1 + \frac{-1st^{2}v^{2} + 2t^{\frac{1}{4}}v + s^{2}t^{\frac{1}{4}}}{45v^{3}}) + (\frac{15t^{2}v^{2} - 2t^{\frac{1}{4}}v - 3s^{2}t^{\frac{1}{4}}}{45v^{3}})_{\mu} + \dots$$

$$(\frac{st^{\frac{1}{4}}}{15v^{3}})_{\mu}^{2} - (\frac{t^{\frac{1}{4}}}{45v^{3}})_{\mu}^{\frac{1}{4}} + \dots$$

# 3. Multiple Integration.

If  $(X_1,X_2)$  are bivariate normal, the conditional distribution of  $X_2$ , given  $X_1$ , is  $h(\mu_2+b_{21}(X_1-\mu_1), \sigma_{2.1}^2)$ . The approximation procedure given above yields (writing p for the vector of coefficients of P, and with C the matrix whose j,kth element is the coefficient of  $\mu^k$  in the approximation of  $I_j/I_0$ ):

$$E(P(X_2)|X_2 \in (s_2-t_2,s_2+t_2),X_1) = (pc(s_2,t_2,\sigma_{2,1}^2))(\mu_2+b_{21}(X_1-\mu_1))$$
.

This last is of course a polynomial in  $X_1$ , with coefficients

$$\mathbf{q}_{\alpha} := \mathbf{b}_{21}^{\alpha} \sum_{\beta > \alpha} (\mathbf{pc}(\mathbf{s}_{2}, \mathbf{t}_{2}, \sigma_{2,1}^{2}))_{\beta} (_{\alpha}^{\beta}) (\mu_{2} - \mathbf{b}_{21} \mu_{1})^{\beta - \alpha},$$

so that

$$\begin{split} \mathbf{E}(\mathbf{P}(\mathbf{X}_{2}) \big| \mathbf{X}_{2} & \in (\mathbf{s}_{2} - \mathbf{t}_{2}, \mathbf{s}_{2} + \mathbf{t}_{2}), \ \mathbf{X}_{1} \in (\mathbf{s}_{1} - \mathbf{t}_{1}, \mathbf{s}_{1} + \mathbf{t}_{1})) = \\ & (\mathbf{q} \mathbf{C}(\mathbf{s}_{1}, \mathbf{t}_{1}, \sigma_{1}^{2}))(\mu_{1}) . \end{split}$$

Other cases (higher dimensions, multivariate polynomials) are treated similarly, but it will be best to introduce some notation at this point.

If p is a vector of coefficients, then

SUBS(a,b,p) is a vector of coefficients,

$$(\text{SUBS}(\mathbf{a},\mathbf{b},\mathbf{p}))_{\mathbf{j}} := \mathbf{b}^{\mathbf{j}} \sum_{\mathbf{k} \geq \mathbf{j}} \binom{\mathbf{k}}{\mathbf{j}} \mathbf{a}^{\mathbf{k}-\mathbf{j}} .$$

(i.e. SUBS(a,b,p) is the vector of coefficients of  $y^j$  when a+by is substituted for x in p(x).

With the aid of a symbolic computation system such as MIT's MACSYMA, the coefficients of the transformation

$$x^k \rightarrow \sum_{\alpha} c_{k\alpha}^{\alpha}^{\alpha}$$
,

as functions  $c_{k\alpha}(v,s,t)$  may be computed for much larger values of k and  $\alpha$  then would be practical by hand. (A partial table of values  $c_{k\alpha}$  is given in ap endix A of this paper. Card versions consisting of FORTRAN assignment statements " c (I,J) =  $\cdots$  " may be obtained on request from the author.) As subsequent examples show, the approximation of  $X_p^k$  in the box  $I_1 \times I_2 \times \cdots I_p$  involves  $c_{m\alpha}$  for higher values of m than k, depending on the order of the underlying Taylor series. Fortunately, the approximation based on only a few terms is quite accurate, as evidenced by the following example: Since the integral

$$s+t\int_{s-t} -\frac{x^2}{2} dx = e^{-\frac{(s-t)^2}{2}} - e^{-\frac{(s+t)^2}{2}}$$

the Taylor series

$$\frac{I_1(1,0,s,t)}{I_0(1,0,s,t)} \sim S(1 - \frac{t^2}{5} + \frac{2+s^2}{45} t^4)$$

gives

$$I_{O}(1,0,s,t) = \Phi(s+t) - \Phi(s-t) \approx \frac{e^{-\frac{(s-t)^{2}}{2}} - e^{-\frac{(s+t)^{2}}{2}}}{\sqrt{2\pi} s(1 - \frac{t^{2}}{3} + \frac{2+s^{2}}{45} t^{l_{1}})}.$$

The approximate values obtained for  $\Phi(x) = .5 + \Phi(x) - \Phi(0)$  by this method agree with tabled values for four decimal places for 0 < x < 1.1.

 $\underline{INT}(s,t,v,\underline{p})$  is a map from the scalars s,t,v and the vector  $\underline{pC}(s,t,v)$ , i.e. matrix postmultiplication, where we take

$$I_k/I_0 \sim \sum_{\alpha=0}^m c_{\alpha,j}\mu^j$$
 as defining C.

We will also consider <u>INT</u> as a function, so that if <u>INT</u> =  $(a_0, a_1, ..., a_m)$ , then  $(\underline{INT})(x) = a_0 + a_1 X + ... + a_m X^m$ .

With the above notation, we can define multiple integrals in a fairly compact fashion. The multivariate normal distribution may be parameterized by a vector of means  $(\mu_1, \mu_2, \dots, \mu_p)$ , the conditional variances  $(V_1, \dots, V_p)$  defined by  $V_1 = \text{Var}(X_1)$ ,  $V_j = \text{Var}(X_j | X_1, \dots, X_{j-1})$ , j > 1, and the regression coefficients  $b_{21}, b_{32}, b_{31}, \dots, b_{p,p-1}, \dots, b_{p1}$  defined by

$$E(X_{j}|X_{1},...,X_{j-1}) = \sum_{i=1}^{j-1} b_{ji}X_{i}$$
.

One-dimensional integration is done directly:

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$$(p,s,t,\mu,v) \leftarrow (INT(s,t,v,p))(\mu)$$
.

Further integrals (I2, I3,...) in higher dimensions are defined recursively, e.g. in two dimensions:

$$\begin{split} & \text{I2}(\mathbf{s}_{1}, \mathbf{t}_{1}, \mathbf{s}_{2}, \mathbf{t}_{2}, \boldsymbol{\mu}_{1}, \boldsymbol{\mu}_{2}, \mathbf{v}_{1}, \mathbf{v}_{2}, \mathbf{b}_{21}, \overset{p}{\sim}) := \\ & \qquad \qquad (\text{INT}(\mathbf{s}_{1}, \mathbf{t}_{1}, \mathbf{v}_{1}, \text{SUBS})\boldsymbol{\mu}_{2} - \mathbf{b}_{21}\boldsymbol{\mu}_{1}, \mathbf{b}_{21}, \text{INT}(\mathbf{s}_{2}, \mathbf{t}_{2}, \mathbf{v}_{2}, \overset{p}{p})))) \ , \end{split}$$

and from this we obtain the two-dimensional integral as

$$({\tt I2}({\tt s}_1,{\tt t}_1,{\tt s}_2,{\tt t}_2,{\tt \mu}_1,{\tt \mu}_2,{\tt v}_1,{\tt v}_2,{\tt b}_{21},{\tt p}))({\tt \mu}_1)\ .$$

In general, IX = INT(...SUBS(...I(K-1))). We have restricted attention here to sets of the form  $I_1 \times I_2 \times \cdots I_p$  where the  $I_j$  are intervals. For more general sets A, a similar scheme would work provided the section of  $X_i \in A$  given  $X_1 \cdots X_{i-1}$  is an interval with endpoints which are polynomials in  $X_1, \dots, X_{i-1}$ .

# 4. A Numeric Implementation.

The following example shows how the three-dimensional conditional expectation approximation may be defined, using only vector and matrix arithmetic with numeric arguments. The polynomial to be approximated is presumed to be of degree at most two, and perhaps bivariate in  $X_2$  and  $X_3$ . (This case is general enough for all the means, squares, and products  $EX_1, EX_2, \ldots EX_1^2, EX_2^2, \ldots EX_1X_2, \ldots EX_2X_3$ ). In fact, we describe here the approximate values of  $E(X^3)$ ,  $E(X_2^2)$ ,  $E(X_2X_3)$ , since other expectations reduce, by interchange of variables, to these three. The approximation used will include powers of t up to  $t^4$  inclusive, i.e.

$$\text{Ex}^{j} \approx \sum_{\alpha=0}^{3} c_{j\alpha}(s,t,v)_{\mu}^{\alpha}$$
.

The functions  $\{c_{j\alpha}\}_{\alpha=0}^3$  6 may be found in appendix A). To approximate  $E(X_3)$ , the integration on  $X_3$  is first carried out, yielding

$$\begin{split} &\sum_{\alpha=0}^{3} \ c_{1\alpha}(s_{3},t_{3},v_{3})(\mu_{3}-\mu_{1}b_{31}-\mu_{2}b_{32}+X_{1}b_{31}+X_{2}b_{32})^{\alpha} \\ &= \sum_{\alpha=0}^{3} \ c_{1\alpha}(s_{3},t_{3},v_{3}) \sum_{\beta=0}^{\alpha} \ \sum_{\gamma=0}^{\alpha-\beta} \ (^{\alpha}_{\beta\gamma}) \ (b_{31}X_{1})^{\beta} (b_{32}X_{2})^{\gamma} (\mu_{3}-\mu_{1}b_{31}-\mu_{2}b_{32})^{\alpha-(\beta+\gamma)} \\ &\sum_{\alpha=0}^{3} \ \sum_{\beta=0}^{3-\alpha} \ (\sum_{\gamma=\alpha+\beta}^{3} \ (^{\gamma}_{\alpha,\beta}) c_{1\gamma}(s_{3},t_{3},v_{3}) (\mu_{3}-\mu_{1}b_{31}-\mu_{2}b_{32})^{\gamma-(\alpha+\beta)}) b_{31}^{\alpha} b_{32}^{\beta} \cdot X_{1}^{\alpha} X_{2}^{\beta} \ . \end{split}$$

The above expression is of the form  $\sum_{\alpha=0}^{3}\sum_{\beta=0}^{3-\alpha}r_{\alpha\beta}X_{1}^{\alpha}X_{2}^{\beta},$  where

$$\mathbf{r}_{\alpha\beta} = \mathbf{b}_{31}^{\alpha} \mathbf{b}_{32}^{\beta} \sum_{\gamma=\alpha+\beta}^{3} {\gamma \choose \alpha,\beta} \mathbf{c}_{1\gamma} (\mathbf{s}_{3},\mathbf{t}_{3},\mathbf{v}_{3}) \cdot (\mu_{3} - \mu_{1} \mathbf{b}_{31} - \mu_{2} \mathbf{b}_{32})^{\gamma - (\alpha+\beta)} .$$

Integration on  $X_2$  then yields:

$$\begin{split} \sum_{\alpha=0}^{3} \sum_{\beta=0}^{3-\alpha} & \mathbf{r}_{\alpha\beta} \ \mathbf{x}_{1}^{\alpha} \ (\sum_{\gamma=0}^{3} \mathbf{c}_{\beta\gamma}(\mathbf{s}_{2},\mathbf{t}_{2},\mathbf{v}_{2}) \cdot (\mathbf{\mu}_{2}-\mathbf{\mu}_{1}\mathbf{b}_{21}+\mathbf{b}_{21}\mathbf{x}_{1})^{\gamma}) \\ &= \sum_{\alpha=0}^{6} \mathbf{s}_{\alpha}\mathbf{x}_{1}^{\alpha} \ , \quad \text{where} \quad \mathbf{s}_{\alpha} = \sum_{\delta=(\alpha-2) \neq 0}^{\alpha} \sum_{\beta=0}^{3-\delta} \mathbf{r}_{\delta\beta} \sum_{\gamma=\alpha-\delta}^{3} (\mathbf{x}_{1}^{\gamma})\mathbf{b}_{21}^{\alpha-\delta} \cdot \\ & \mathbf{c}_{\beta\gamma}(\mathbf{s}_{2},\mathbf{t}_{2},\mathbf{v}_{2}) \cdot (\mathbf{\mu}_{2}-\mathbf{\mu}_{1}\mathbf{b}_{21})^{\gamma-(\alpha-\delta)} \ , \end{split}$$

and the final form is

$$\sum_{\alpha=0}^6 s_\alpha \sum_{\beta=0}^3 c_{\alpha\beta}(\boldsymbol{s}_1,\boldsymbol{t}_1,\boldsymbol{v}_1) \mu_1^\beta \ .$$

The approximation of  $\chi_3^2$  is carried out in the same fashion:

$$\begin{split} \text{EX}_{3}^{2} &= \sum_{\alpha=0}^{6} \, \mathbf{s}_{\alpha}^{i} \, \sum_{\beta=0}^{3} \, \mathbf{c}_{\alpha\beta}(\mathbf{s}_{1}, \mathbf{t}_{1}, \mathbf{v}_{1}) \mu_{1}^{\beta} \, , \\ \\ \mathbf{s}_{\alpha}^{i} &= \sum_{\delta=(\alpha-3)}^{6} \sum_{\mathbf{v}} \, \sum_{\beta=0}^{3-\delta} \, \mathbf{r}_{\delta\beta}^{i} \, \sum_{\gamma=\alpha-\delta}^{3} (\gamma_{\alpha-\delta}^{\gamma}) \mathbf{b}_{21}^{\alpha-\delta} \mathbf{c}_{\beta\gamma}(\mathbf{s}_{2}, \mathbf{t}_{2}, \mathbf{v}_{2}) \, . \\ \\ &\qquad \qquad (\mu_{2} - \mu_{1} \mathbf{b}_{21})^{\gamma - (\alpha - \delta)} \, , \end{split}$$

$$\mathbf{r}'_{\alpha\beta} = b_{31}^{\alpha} b_{32}^{\beta} \sum_{\gamma=\alpha+\beta}^{3} {\gamma \choose \alpha\beta} c_{2\gamma} (s_3, t_3, v_3) (\mu_3 - \mu_1 b_{31} - \mu_2 b_{32})^{\gamma - (\alpha+\beta)}$$
.

(Note that the only change is the substitution of  $c_{2\gamma}(s_3,t_3,v_3)$  for  $c_{1\gamma}(s_3,t_3,v_3)$ ).

To approximate  $X_2X_3$ , the first approximate integration is exactly as in approximating  $X_3$ , but the resulting sum is

$$X_2(\sum_{\alpha=0}^{3}\sum_{\beta=0}^{3-\alpha}r_{\alpha\beta}X_1^{\alpha}X_2^{\beta})$$
,

i.e. we have  $r_{\alpha\beta}^{"} = r_{\alpha(\beta-1)}$ ,  $\beta = 1,...,4-\alpha$ . A second integration gives

$$\sum_{\alpha=0}^{3} \sum_{\beta=1}^{4-\alpha} r_{\alpha\beta}^{"} x_{1}^{\alpha} (\sum_{\gamma=0}^{3} c_{\beta\gamma}(s_{2}, t_{2}, v_{2}) (\mu_{2} - b_{21} \mu_{1} + b_{21} X_{1})^{\gamma}),$$

so the final result is

$$\sum_{\alpha=0}^{6} s_{\alpha}^{"} \sum_{\beta=0}^{3} c_{\alpha\beta}(s_{1}, t_{1}, v_{1}) \mu_{1}^{\beta} ,$$

where

$$\mathbf{s}_{\alpha}^{"} = \sum_{\delta=(\alpha-3)}^{\alpha} \sum_{\mathbf{v}=0}^{\mu-\delta} \mathbf{r}_{\delta\beta} \sum_{\gamma=\alpha-\delta}^{3} (\gamma) \mathbf{b}_{21}^{\alpha-\delta} \mathbf{c}_{\beta\gamma} (\mathbf{s}_{2}, \mathbf{t}_{2}, \mathbf{v}_{1}) (\mu_{2} - \mu_{1} \mathbf{b}_{21})^{\gamma-(\alpha-\delta)}.$$

# Higher Order Approximation, Higher Dimensions.

In general, if EX =  $\sum_{\alpha=0}^{N} c_{1\alpha}^{\alpha}^{\alpha}$  is taken as the basic approximation, the three-dimensional integral will be of the form

$$\sum_{\alpha=0}^{2N} s_{\alpha} \sum_{\beta=0}^{N} c_{\alpha\beta}(s_{1},t_{1},v_{1})_{\mu_{1}}^{\beta},$$

where

$$\mathbf{s}_{\alpha} = \sum_{\alpha} \sum_{\beta} \mathbf{s}_{\alpha\beta}^{(2)}, \quad \mathbf{s}_{\alpha\beta}^{(2)} = \sum_{\alpha\beta} \sum_{\alpha\beta\gamma} \mathbf{s}_{\alpha\beta\gamma}^{(3)}, \dots, \quad \mathbf{s}_{\alpha_{1}\alpha_{2}}^{(p-1)}$$

Thus, increasing accuracy is computationally expensive, and increasing dimension more so, but nonetheless the process should compare favorably with numerical integration, as borne out by some preliminary comparisons, where Romberg integration was about 10x slower.

Appendix A: Values of cjk

Interval (s-t,s+t)  

$$X \sim N(\mu, \sigma^{2})$$

$$W = \frac{1}{\sigma^{2}} = \frac{1}{v}$$

j = 1

**k**≈0

 $-\frac{s(t^2w(t^2w(w(2t^2(s^2w(s^2w+4)+1)-21s^2)-42)+315-945)}{945}$ 

k=l

 $\frac{t^2w(w(2t^2(s^2w(5s^2w+12)+1)-63s^2)-42+315)}{945}$ 

k=2

 $= \frac{st^{4}w^{3}(4t^{2}w(5s^{2}w+6)-63)}{945}$ 

k=3

 $\frac{t^4w^3(4t^2w(5s^2w+2)-21)}{945}$ 

k=4

 $-\frac{2st^6w^5}{189}$ 

k=5

2t<sup>6</sup>w<sup>5</sup>

**k**=0

- 
$$(t^2(2w(t^2(w(t^2(2s^2w(s^2w+5)+2)-1)-21s^2(s^2w+3))+21)+315s^2)-315)$$
  
-945s<sup>2</sup>)/945

k=1

$$\frac{2st^2w(t^2w(w(2t^2(s^2w+3)(5s^2w+1)-63s^2)-84)+315)}{945}$$

**k**=2

$$-\frac{2t^{4}w^{2}(w(2t^{2}(2s^{2}w(5s^{2}w+9)+1)-63s^{2})-21)}{945}$$

k=3

$$\frac{2st^{4}w^{3}(4t^{2}w(5s^{2}w+4)-21)}{945}$$

k=4

$$-\frac{4t^{6}w^{4}(5s^{2}w+1)}{945}$$

k=5

**k=**0

 $-s(t^{2}(w(t^{2}(s^{2}w(2s^{2}w(s^{2}w+6)+3)-14)-21s^{2}(s^{2}w+4))$   $+105)+315s^{2})-315)-315s^{2})/315$ 

k=1  $t^2w(t^2(w(t^2(s^2w(10s^2w(s^2w+4)+1)-12)-63s^2(s^2w+2))+63)$   $+ 315s^2)/315$ 

 $k\approx 2$   $= \frac{st^{\frac{1}{4}}w^{2}(w(t^{2}(2s^{2}w+5)(10s^{2}w-1)-63s^{2})-42)}{315}$ 

k≈3

 $\frac{t^4w^3(t^2(4s^2w(5s^2w+6)-3)-21s^2)}{315}$ 

k=4

 $-\frac{2st^{6}w^{4}(5s^{2}w+2)}{315}$ 

k≈5

 $\frac{2s^2t^6w^5}{315}$ 

k=0

 $-(888t^6w^5+5686t^6w^4-48^4t^6w^3-8486t^4w^3-1928^2t^6w^2)$ 

 $-420s^{1}t^{1}w^{2}+36t^{6}w+1008s^{2}t^{1}w+1260s^{1}t^{2}w-189t^{1}-1890s^{2}t^{2}$ 

- 945s<sup>4</sup>)/945

k=1

 $4st^2w(t^2(w(t^2(s^2w(2s^2w(5s^2w+24)-13)-54)$   $-21s^2(3s^2w+8))+189)+315s^2)/915$ 

k=2

 $-\frac{4t^{1/4}w^{2}(t^{2}(s^{2}w(20s^{2}w(s^{2}w+3)-21)-9)-63s^{2}(s^{2}w+1))}{945}$ 

k=3

 $\frac{4st^{4}w^{3}(t^{2}(4s^{2}w(5s^{2}w+8)-9)-21s^{2})}{945}$ 

k=4

 $-\frac{8s^2t^6w^4(5s^2w+3)}{945}$ 

k=5

 $\frac{8s^3t^6w^5}{945}$ 

k=0

 $-s(2s^8t^6w^5+16s^6t^6w^4-8s^4t^6w^3-21s^6t^4w^3-112s^2t^6w^2$ 

 $-126s^{4}t^{4}w^{2}+63t^{6}w+462s^{2}t^{4}w+315s^{4}t^{2}w-189t^{4}-630s^{2}t^{2}$ 

- 189s<sup>1</sup>)/189

k=1

 $t^{2}w(10s^{8}t^{1}w^{1}+56s^{6}t^{1}w^{3}-36s^{1}t^{1}w^{2}-63s^{6}t^{2}w^{2}-144s^{2}t^{1}w$   $-210s^{1}t^{2}w+27t^{1}+378s^{2}t^{2}+315s^{1})/189$ 

k=2

 $-\frac{st^{4}w^{2}(2t^{2}(s^{2}w(2s^{2}w(5s^{2}w+18)-23)-18)-21s^{2}(3s^{2}w+4))}{189}$ 

k=3

 $\frac{s^2t^4w^3(2t^2(10s^2w(s^2w+2)-9)-21s^2)}{189}$ 

k\_l

 $-\frac{2s^{3}t^{6}w^{4}(5s^{2}w+4)}{189}$ 

k=5

 $\frac{2s^4t^6w^5}{189}$ 

k=0

- $-(4s^{10}t^{6}w^{5}+36s^{8}t^{6}w^{4}-36s^{6}t^{6}w^{3}-42s^{8}t^{4}w^{3}-430s^{4}t^{6}w^{2}$
- $-294s^{6}t^{4}w^{2}+450s^{2}t^{6}w+1470s^{4}t^{4}w+630s^{6}t^{2}w-45t^{6}-945s^{2}t^{4}$
- 1575s<sup>4</sup>t<sup>2</sup>-315s<sup>6</sup>)/315

k=l

 $2st^{2}w(10s^{8}t^{\frac{1}{4}}w^{\frac{1}{4}}+64s^{6}t^{\frac{1}{4}}w^{3}-68s^{\frac{1}{4}}t^{\frac{1}{4}}w^{2}-63s^{6}t^{2}w^{2}-300s^{2}t^{\frac{1}{4}}w$  $-252s^{\frac{1}{4}}t^{2}w+135t^{\frac{1}{4}}+630s^{2}t^{2}+315s^{\frac{1}{4}})/315$ 

k=2

 $-2s^2t^4w^2(2t^2(2s^2w(s^2w+5)(5s^2w-4)-45)-21s^2(3s^2w+5))/315$ 

k≈3

 $\frac{2s^3t^{\frac{1}{4}}w^3(2t^2(2s^2w(5s^2w+12)-15)-21s^2)}{315}$ 

k=4

 $\frac{4s^{4}t^{6}w^{4}(s^{2}w+1)}{63}$ 

k=5

 $\frac{4s^5t^6w^5}{315}$ 

- $s(2s^{10}t^{6}w^{5}+20s^{8}t^{6}w^{4}-31s^{6}t^{6}w^{3}-21s^{8}t^{4}w^{3}-366s^{4}t^{6}w^{2}$
- $-168s^6t^4w^2+585s^2t^6w+1071s^4t^4w+315s^6t^2w-135t^6$
- $-9458^{2}t^{4}-9458^{4}t^{2}-1358^{6})/135$

k=1

 $s^{2}t^{2}w(10s^{8}t^{4}w^{4}+72s^{6}t^{4}w^{3}-109s^{4}t^{4}w^{2}-63s^{6}t^{2}w^{2}-540s^{2}t^{4}w$ -  $294s^{4}t^{2}w+405t^{4}+945s^{2}t^{2}+315s^{4})/315$ 

k=2

 $\frac{s^3t^4w^2(t^2(s^2w(4s^2w(5s^2w+24)-123)-180)-63s^2(s^2w+2))}{135}$ 

k=3

 $\frac{s^{\frac{1}{4}}t^{\frac{1}{4}}w^{3}(t^{2}(\frac{1}{4}s^{2}w(5s^{2}w+1\frac{1}{4})-\frac{1}{4}5)-21s^{2})}{135}$ 

k=3

 $-\frac{2s^5t^6w^4(5s^2w+6)}{135}$ 

k=5

 $\frac{2s^6t^6y^5}{135}$ 

**k**=0

 $-s^2(16s^{10}t^6w^5+176s^8t^6w^4-376s^6t^6w^3-168s^8t^4w^3$ 

 $-4592s^{4}t^{6}w^{2}-1512s^{6}t^{4}w^{2}+10080s^{2}t^{6}w+11760s^{4}t^{4}w$ 

 $+2520s^{6}t^{2}w-3780t^{6}-13230s^{2}t^{4}-8820s^{4}t^{2}-945s^{6})/945$ 

k=1

 $8s^3t^2w(10s^8t^4w^4+80s^6t^4w^3-159s^4t^4w^2-68s^6t^2w^2$ 

ので、10mmでは、10mm

 $-882s^2t^4w-336s^4t^2w+945t^4+1323s^2t^2+315s^4)/945$ 

k=2

 $8s^{4}t^{4}w^{2}(t^{2}(s^{2}w(4s^{2}w(5s^{2}w+27)-175)-315)-21s^{2}(3s^{2}w+7))/945$ 

k=3

 $\frac{8s^5t^4w^3(t^2(4s^2w(5s^2w+16)-63)-21s^2}{945}$ 

k=4

 $-\frac{16s^6t^6w^4(5s^2w+7)}{945}$ 

k=5

 $\frac{16s^7t^6w^5}{945}$ 

- $-s^{3}(2s^{10}t^{6}w^{5}+24s^{8}t^{6}w^{4}-66s^{6}t^{6}w^{3}-21s^{8}t^{4}w^{3}-848s^{4}t^{6}w^{2}$
- $= 210s^6t^4w^2 + 2394s^2t^6w + 1932s^4t^4w + 315s^6t^2w 1260t^2$
- $-2646s^2t^4-1260s^4t^2-105s^6)/105$

k≈l

 $s^{4}t^{2}w(10s^{8}t^{4}w^{4}+88s^{6}t^{4}w^{3}-218s^{4}t^{4}w^{2}-63s^{6}t^{2}w^{2}-1344s^{2}t^{4}w$ 

 $-387s^{4}t^{2}w+1890t^{4}+1764s^{2}t^{2}+315s^{4})/105$ 

k=2

 $= \frac{s^5 t^{1/4} w^2 (4 t^2 (s^2 w (5 s^2 w (s^2 w + 6) - 59) - 126) - 21 s^2 (3 s^2 w + 8))}{105}$ 

k=3

 $\frac{s^6t^4w^3(4t^2(s^2w(5s^2w+18)-21)-21s^2)}{105}$ 

k=4

 $-\frac{2s^7t^6w^4(5s^2w+8)}{105}$ 

k=5

 $\frac{2\mathbf{s}^{8}\mathbf{t}^{6}\mathbf{v}^{5}}{105}$ 

k=0

- $-s^{4}(4s^{10}t^{6}w^{5}+52s^{8}t^{6}w^{4}-176s^{6}t^{6}w^{3}-42s^{8}t^{4}w^{3}-2394s^{4}t^{6}w^{2}$
- $-462s^6t^4w^2+8316s^2t^6w+4914s^4t^4w+630s^6t^2w-5670t^6$
- $-7938s^2t^4-2835s^4t^2-189s^6)/189$

k=1  $2s^{5}t^{2}w(10s^{8}t^{4}w^{4}+96s^{6}t^{4}w^{3}-286s^{4}t^{4}w^{2}-63s^{6}t^{2}w^{2}$ 

 $-1944s^{2}t^{4}w-420s^{4}t^{2}w+3402t^{4}+2268s^{2}t^{2}+315s^{4})/189$ 

**k=**2

 $2s^6t^4w^2(2t^2(s^2w(2s^2w(5s^2w+33)-153)-378)-63s^2(s^2w+3))/189$ 

k=3

 $\frac{2s^7t^{\frac{1}{4}}w^{\frac{3}{2}}(\frac{4t^2(5s^2w(s^2w+\frac{1}{4})-27)-21s^2})}{189}$ 

k=4

 $-\frac{4s^8t^6w^4(5s^2w+9)}{189}$ 

k=5

 $\frac{4s^9t^6w^5}{189}$ 

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#### TECHNICAL REPORTS

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